# Design and Implementation of Speech Recognition Systems 

Spring 2011

Class 8: HMMs<br>14 Feb 2011

## Recap: Generalized Templates



- A set of "states"
- A distance function associated with each state
- A set of transitions
- Transition-specific penalties


## Recap: HMMs



- Identical to generalized templates in principle
- "Distance" functions at states replaced by "probability distribution function" for state
- Transition "penalties" replaced by transition probabilities
- Maximize probability of observation
- Instead of minimizing cost
- The entire structure may be viewed as one generalization of the DTW models we have discussed thus far


## The HMM Process

- The HMM models the process underlying the observations as going through a number of states
- E.g., to produce the sound "W", it first goes through a state where it produces the sound "UH", then goes into a state where it transitions from "UH" to "AH", and finally to a state where it produced "AH"

- The true underlying process is the vocal tract here
- Which roughly goes from the configuration for "UH" to the configuration for " AH "


## HMMs are abstractions

- The states are not directly observed
- Here states of the process are analogous to configurations of the vocal tract that produces the signal
- We only hear the speech; we do not see the vocal tract
- i.e. the states are hidden
- The interpretation of states is not always obvious
- The vocal tract actually goes through a continuum of configurations
- The model represents all of these using only a fixed number of states
- The model abstracts the process that generates the data
- The system goes through a finite number of states
- When in any state it can either remain at that state, or go to another with some probability
- When at any states it generates observations according to a distribution associated with that state


## Hidden Markov Models



- A Hidden Markov Model consists of two components
- A state/transition backbone that specifies how many states there are, and how they can follow one another
- A set of probability distributions, one for each state, which specifies the distribution of all vectors in that state


Markov chain

- This can be factored into two separate probabilistic entities
- A probabilistic Markov chain with states and transitions
- A set of data probability distributions, associated with the states


## Equivalence to DTW templates

- HMM - inference equivalent to DTW modified to use a probabilistic function, for the local node or edge "costs" in the trellis
- Edges have transition probabilities
- Nodes have output or observation probabilities
- They provide the probability of the observed input
- The output probability may be a Gaussian
- Goal is to find the template with highest probability of matching the input
- Probability values associated with transitions and edges are called likelihoods


## Likelihoods and Cost: Transition



Markov chain

DTW Template

- Transitions in the HMM have associated probabilities
$-\mathrm{P}_{11}, \mathrm{P}_{12}$ etc
- They can be converted to "scores" through a logarithm
- $\mathrm{T}_{11}=\log \left(\mathrm{P}_{11}\right)$
- Or to "costs" through a negative logarithm
$-\mathrm{T}_{11}=-\log \left(\mathrm{P}_{11}\right)$


## Likelihoods and Cost: Nodes



- States in the HMM have probability distributions associated with them
- E.g Gaussians
- Whose means and variances have been obtained from the segments associated with the node
- Nodes in the trellis have a probabilities associated with them
$-\mathrm{P}_{\mathrm{i}}(\mathrm{O})$
- $i$ is the "state" / template node
- $O$ is the observation associated with any node in the trellis
- Node probabilities may be converted to:
- Scores: $\mathrm{N}_{\mathrm{i}}(\mathrm{O})=\log \left(\mathrm{P}_{\mathrm{i}}(\mathrm{O})\right)$
- Or Costs: $\mathrm{N}_{\mathrm{i}}(\mathrm{O})=-\log \left(\mathrm{P}_{\mathrm{i}}(\mathrm{O})\right)$


## Log Likelihoods



- Use probabilities or likelihoods instead of cost
- Scores combines multiplicatively along a path
$-\quad$ Path Score $=\mathrm{P}_{1}\left(\mathrm{O}_{1}\right) \cdot \mathrm{P}_{11} \cdot \mathrm{P}_{1}\left(\mathrm{O}_{2}\right) \cdot \mathrm{P}_{12} \cdot \mathrm{P}_{2}\left(\mathrm{O}_{3}\right) \cdot \mathrm{P}_{22} \cdot \mathrm{P}_{2}\left(\mathrm{O}_{4}\right) \cdot \mathrm{P}_{23} \cdot \mathrm{P}_{3}\left(\mathrm{O}_{5}\right) \cdot \mathrm{P}_{23}$
$\square \quad$ Alternately use $\log$ probabilities as scores: $\mathrm{N}_{\mathrm{i}}(\mathrm{O})=\log \left(\mathrm{P}_{\mathrm{i}}(\mathrm{O})\right), \mathrm{T}_{11}=\log \left(\mathrm{P}_{11}\right)$
- Scores add as in DTW
- Path Score $=\mathrm{N}_{1}\left(\mathrm{O}_{1}\right)+\mathrm{T}_{11}+\mathrm{N}_{1}\left(\mathrm{O}_{2}\right)+\mathrm{T}_{12}+\mathrm{N}_{2}\left(\mathrm{O}_{3}\right)+\mathrm{T}_{22}+\mathrm{N}_{2}\left(\mathrm{O}_{4}\right)+\mathrm{T}_{23}+\mathrm{N}_{3}\left(\mathrm{O}_{5}\right)+\mathrm{T}_{23}$
- Replace all "Min" operations in DTW by "Max"
- Alternately use negative $\log$ probabilities as cost: $\mathrm{N}_{\mathrm{i}}(\mathrm{O})=\log \left(\mathrm{P}_{\mathrm{i}}(\mathrm{O})\right), \mathrm{T}_{11}=-\log \left(\mathrm{P}_{11}\right)$
- Cost adds as in DTW
- Computation remains identical to DTW (with edge costs factored in)


## HMM as a statistical model

- An HMM is a statistical model for a time-varying process
- The process is always in one of a countable number of
- When the process visits in any state, it generates an observation by a random draw from a distribution associated with that state
- The process constantly moves from state to state. The probability that the process will move to any state is determined solely by the current state
- i.e. the dynamics of the process are Markovian
- The entire model represents a probability distribution over the sequence of observations
- It has a specific probability of generating any particular sequence
- The probabilities of all possible observation sequences sums to 1


## How an HMM models a process



## HMM Parameters

- The topology of the HMM
- No. of states and allowed transitions
- E.g. here we have 3 states and cannot go from the blue state to the red
- The transition probabilities
- Often represented as a matrix as here
$-T_{i j}$ is the probability that when in state i , the process will move to j

- The probability of being at a particular state at the first instant
- The state output distributions


## HMM state output distributions

- The state output distribution represents the distribution of data produced from any state
- In the previous lecture we assume the state output distribution to be Gaussian
- Albeit largely in a DTW context

$$
P(v)=\operatorname{Gaussian}(v ; m, C)=\frac{1}{\sqrt{2 \pi|C|}} e^{-0.5(v-m)^{T} C^{-1}(v-m)}
$$

- In reality, the distribution of vectors for any state need not be Gaussian
- In the most general case it can be arbitrarily complex
- The Gaussian is only a coarse representation of this distribution
- If we model the output distributions of states better, we can expect the model to be a better representation of the data


## Node Score: The Gaussian Distribution

- What does a Gaussian distribution look like?
- For a single (scalar) variable, it is a bell-shaped curve representing the density of data around the mean
- Example:


| Four different scalar Gaussian <br> distributions, with different means and <br> variances <br> The mean is represented by $\mu$, and <br> variance by $\sigma^{2}$ <br> $\mu$ and $\sigma$ are the parameters of the <br> Gaussian distribution <br> (Taken from Wikipedia) |
| :--- |

Four different scalar Gaussian distributions, with different means and variances

The mean is represented by $\mu$, and variance by $\sigma^{2}$
$\mu$ and $\sigma$ are the parameters of the Gaussian distribution
(Taken from Wikipedia)

## The Scalar Gaussian Function

- The Gaussian density function (the bell curve) is:

$$
p(x)=\frac{1}{\sqrt{2 \pi} \sigma} e^{\mathrm{L} \frac{1}{2 \sigma^{2}}(x-\mu)^{2}}
$$

- $p(x)$ is the density function of the variable $x$, with mean $\mu$ and variance $\sigma^{2}$
- The attraction of the Gaussian function (regardless of how appropriate it is!) comes from how easily the mean and variance can be estimated from sample data $x_{1}, x_{2}, x_{3} \ldots x_{N}$
- $\mu=\Sigma_{i} x_{i} / N$
$-\sigma^{2}=\sum_{i}\left(x_{i}-\mu\right)^{2} / N=\sum_{i}\left(x_{i}^{2}-\mu^{2}\right) / N$


## The 2-D Gaussian Distribution

- Speech data are not scalar values, but vectors!
- Needs multi-variate (multi-dimensional) Gaussians
- Figure: A Gaussian for 2-D data
- Shown as a 3-D plot
- Distributions for higher dimensions are tough to visualize!


## The Multidimensional Gaussian Distribution

- Instead of variance, the multidimensional Gaussian has a covariance matrix
- The multi-dimensional Gaussian distribution of a vector variable $x$ with mean $\mu$ and covariance $\Sigma$ is given by:

$$
f(x)=(2 \pi)^{-N / 2} \operatorname{det}(\Sigma)^{-1 / 2} \exp \left(-\frac{1}{2}(x-\mu)^{T} \Sigma^{-1}(x-\mu)\right)
$$

- where $N$ is the vector dimensionality, and det is the determinant function
- The complexity in a full multi-dimensional Gaussian distribution comes from the covariance matrix, which accounts for dependencies between the dimensions


## The Diagonal Covariance Matrix

Full covariance:
all elements are
non-zero


Diagonal covariance:
off-diagonal elements are zero


- In speech recognition, we frequently assume that the feature vector dimensions are all independent of each other
- Result: The covariance matrix is reduced to a diagonal form
- The exponential term becomes, simply:

$$
\left(\Sigma_{\mathrm{i}}\left(x_{i}-\mu_{i}\right)^{2} / \sigma_{i}^{2}\right) / 2, i \text { going over all vector dimensions }
$$

- The determinant of the diagonal $\Sigma$ matrix is easy to compute
- Further, each $\sigma_{i}^{2}$ (the $i$-th digonal element in the covariance matrix) is easily estimated from $x_{i}$ and $\mu_{i}$ like a scalar


## Gaussian Mixtures

- A Gaussian Mixture is literally a mixture of Gaussians. It is a weighted combination of several Gaussian distributions

$$
P(v)=\sum_{i=0}^{K-1} w_{i} \operatorname{Gaussian}\left(v ; \mu_{i}, C_{i}\right)
$$

- $v$ is any data vector. $P(v)$ is the probability given to that vector by the Gaussian mixture
- K is the number of Gaussians being mixed
- $\mathrm{w}_{\mathrm{i}}$ is the mixture weight of the $\mathrm{i}^{\text {th }}$ Gaussian. $\mu_{\mathrm{i}}$ is its mean and $\mathrm{C}_{\mathrm{i}}$ is its covariance
- The Gaussian mixture distribution is also a distribution
- It is positive everywhere.
- The total volume under a Gaussian mixture is 1.0.
- Constraint: the mixture weights $\mathrm{w}_{\mathrm{i}}$ must all be positive and sum to 1


## Gaussian Mixtures

- A Gaussian mixture can represent data distributions far better than a simple Gaussian
- The two panels show the histogram of an unknown random variable
- The first panel shows how it is modeled by a simple Gaussian
- The second panel models the histogram by a mixture of two Gaussians
- Caveat: It is hard to know the optimal number of Gaussians in a mixture distribution for any random variable



## Generating an observation from a Gaussian mixture state distribution

First draw the identity of the Gaussian from the a priori probability distribution of Gaussians (mixture weights)

Then draw a vector from the selected Gaussian


## HMMs with Gaussian mixture state distributions

- The parameters of an HMM with Gaussian mixture state distributions are:
- $\pi$ the set of initial state probabilities for all states
- $\boldsymbol{T}$ the matrix of transition probabilities
- A Gaussian mixture distribution for every state in the HMM. The Gaussian mixture for the $i^{\text {th }}$ state is characterized by
- $\mathrm{K}_{\mathrm{i}}$, the number of Gaussians in the mixture for the $\mathrm{i}^{\text {th }}$ state
- The set of mixture weights $\mathrm{w}_{\mathrm{i}, \mathrm{j}} 0<\mathrm{j}<\mathrm{K}_{\mathrm{i}}$
- The set of Gaussian means $\mu_{\mathrm{i}, \mathrm{j}} 0<\mathrm{j}<\mathrm{K}_{\mathrm{i}}$
- The set of Covariance matrices $\mathrm{C}_{\mathrm{i}, \mathrm{j}} 0<\mathrm{j}<\mathrm{K}_{\mathrm{i}}$


## Three Basic HMM Problems

- Given an HMM:
- What is the probability that it will generate a specific observation sequence
- Given a observation sequence, how do we determine which observation was generated from which state
- The state segmentation problem
- How do we learn the parameters of the HMM from observation sequences

Computing the Probability of an Observation Sequence

- Two aspects to producing the observation:
- Progressing through a sequence of states
- Producing observations from these states


## Progressing through states

HMM assumed to be generating data

state sequence


- The process begins at some state (red) here
- From that state, it makes an allowed transition
- To arrive at the same or any other state
- From that state it makes another allowed transition
- And so on

Probability that the HMM will follow a particular state sequence

$$
P\left(s_{1}, s_{2}, s_{3}, \ldots\right)=P\left(s_{1}\right) P\left(s_{2} \mid s_{1}\right) P\left(s_{3} \mid s_{2}\right) \ldots
$$

- $P\left(s_{l}\right)$ is the probability that the process will initially be in state $s_{l}$
- $P\left(s_{i} \mid s_{j}\right)$ is the transition probability of moving to state $s_{i}$ at the next time instant when the system is currently in $s_{j}$
- Also denoted by $\mathrm{P}_{\mathrm{ij}}$ earlier
- Related to edge scores in DTW as $\mathrm{T}_{\mathrm{ij}}=-\log \left(P\left(s_{i} \mid s_{j}\right)\right)$


## Generating Observations from States



- At each time it generates an observation from the state it is in at that time

Probability that the HMM will generate a particular observation sequence given a state sequence (state sequence known)

$$
P\left(o_{1}, o_{2}, o_{3}, \ldots \mid s_{1}, s_{2}, s_{3}, \ldots\right)=P\left(o_{0} \mid s_{1}\right) P\left(o_{2} \mid s_{2}\right) P\left(\left.o_{3}\right|_{3}\right) \ldots
$$

Computed from the Gaussian or Gaussian mixture for state $s_{1}$

- $P\left(o_{i} \mid s_{i}\right)$ is the probability of generating observation $o_{i}$ when the system is in state $s_{i}$
- Related to node scores in DTW trellis as:

$$
N_{i}(\mathrm{O})=-\log \left(P\left(o_{i} \mid s_{i}\right)\right)
$$

Progressing through States and Producing Observations

HMM assumed to be generating data
state


- At each time it produces an observation and makes a transition

Probability that the HMM will generate a particular state sequence and, from it, generate a particular observation sequence

$$
\begin{aligned}
& P\left(o_{1}, o_{2}, o_{3}, \ldots, s_{1}, s_{2}, s_{3}, \ldots\right)= \\
& P\left(o_{1}, o_{2}, o_{3}, \ldots \mid s_{1}, s_{2}, s_{3}, \ldots\right) P\left(s_{1}, s_{2}, s_{3}, \ldots\right)= \\
& P\left(o_{1} \mid s_{1}\right) P\left(o_{2} \mid s_{2}\right) P\left(o_{3} \mid s_{3}\right) \ldots P\left(s_{1}\right) P\left(s_{2} \mid s_{1}\right) P\left(s_{3} \mid s_{2}\right) \ldots
\end{aligned}
$$

## Probability of Generating an Observation

 Sequence- If only the observation is known, the precise state sequence followed to produce it is not known
- All possible state sequences must be considered

$$
P\left(o_{1}, o_{2}, o_{3}, \ldots\right)=\sum_{\text {all.possible }} P\left(o_{1}, o_{2}, o_{3}, \ldots, s_{1}, s_{2}, s_{3}, \ldots\right)=
$$

state .sequences
$\sum_{\text {all.posisible }} P\left(o_{1} \mid s_{1}\right) P\left(o_{2} \mid s_{2}\right) P\left(o_{3} \mid s_{3}\right) \ldots P\left(s_{1}\right) P\left(s_{2} \mid s_{1}\right) P\left(s_{3} \mid s_{2}\right) \ldots$ state .sequences

## Computing it Efficiently

- Explicit summing over all state sequences is not efficient
- A very large number of possible state sequences
- For long observation sequences it may be intractable
- Fortunately, we have an efficient algorithm for this: The forward algorithm
- At each time, for each state compute the total probability of all state sequences that generate observations until that time and end at that state


## Illustrative Example



- Consider a generic HMM with 5 states and a "terminating state". We wish to find the probability of the best state sequence for an observation sequence assuming it was generated by this HMM
- $P\left(s_{i}\right)=1$ for state 1 and 0 for others
- The arrows represent transition for which the probability is not 0 . $P\left(s_{i} \mid s_{j}\right)=a_{i j}$
- We sometimes also represent the state output probability of $s_{i}$ as $P\left(o_{t} \mid\right.$ $\left.s_{i}\right)=b_{i}(t)$ for brevity


## Diversion: The HMM Trellis



## Feature vectors (time)

- The trellis is a graphical representation of all possible state sequences through the HMM to produce a given observation
- Analogous to the DTW search graph / trellis
- The Y-axis represents HMM states, X axis represents observations
- Edges in trellis represent valid transitions in the HMM over a single time step
- Every node represents the event of a particular observation being generated from a particular state


## The Forward Algorithm

$$
\alpha_{u}(s, t)=P\left(x_{u, 1}, x_{u, 2}, \ldots, x_{u, t}, \operatorname{state}(t)=s \mid \lambda\right)
$$

- $\alpha_{u}(s, t)$ is the total probability of ALL state sequences that end at state $s$ at time $t$, and all observations until $x_{t}$


## The Forward Algorithm



- $\alpha_{u}(s, t)$ can be recursively computed in terms of $\alpha_{u}\left(s^{\prime}, t^{\prime}\right)$, the forward probabilities at time $\mathrm{t}-1$


## The Forward Algorithm



- In the final observation the alpha at each state gives the probability of all state sequences ending at that state
- The total probability of the observation is the sum of the alpha values at all states

Problem 2: The state segmentation problem

- Given only a sequence of observations, how do we determine which sequence of states was followed in producing it?


## The HMM as a generator



- The process goes through a series of states and produces observations from them


## States are Hidden

HMM assumed to be generating data

state sequence
state


- The observations do not reveal the underlying state


## The state segmentation problem

HMM assumed to be generating data

state
sequence
state distributions observation sequence


- State segmentation: Estimate state sequence given observations


## Estimating the State Sequence

- Any number of state sequences could have been traversed in producing the observation
- In the worst case every state sequence may have produced it
- Solution: Identify the most probable state sequence
- The state sequence for which the probability of progressing through that sequence and gen erating the observation sequence is maximum
- i.e $P\left(o_{1}, o_{2}, o_{3}, \ldots, s_{1}, s_{2}, s_{3}, \ldots\right)$ is maximum


## Estimating the state sequence

- Once again, exhaustive evaluation is impossibly expensive
- But once again a simple dynamic-programming solution is available

$$
\begin{aligned}
& P\left(o_{1}, o_{2}, o_{3}, \ldots, s_{1}, s_{2}, s_{3}, \ldots\right)= \\
& P\left(o_{1} \mid s_{1}\right) P\left(o_{2} \mid s_{2}\right) P\left(o_{3} \mid s_{3}\right) \ldots P\left(s_{1}\right) P\left(s_{2} \mid s_{1}\right) P\left(s_{3} \mid s_{2}\right) \ldots
\end{aligned}
$$

- Needed:
$\arg \max _{s_{1}, s_{2}, s_{3}, \ldots} P\left(o_{1} \mid s_{1}\right) P\left(s_{1}\right) P\left(o_{2} \mid s_{2}\right) P\left(s_{2} \mid s_{1}\right) P\left(o_{3} \mid s_{3}\right) P\left(s_{3} \mid s_{2}\right)$


## Estimating the state sequence

- Once again, exhaustive evaluation is impossibly expensive
- But once again a simple dynamic-programming solution is available

$$
\begin{aligned}
& P\left(o_{1}, o_{2}, o_{3}, \ldots, s_{1}, s_{2}, s_{3}, \ldots\right)= \\
& P\left(o_{1} \mid s_{1}\right) P\left(o_{2} \mid s_{2}\right) P\left(o_{3} \mid s_{3}\right) \ldots P\left(s_{1}\right) P\left(s_{2} \mid s_{1}\right) P\left(s_{3} \mid s_{2}\right) \ldots
\end{aligned}
$$

- Needed:



## The state sequence

- The probability of a state sequence $?, ?, ?, ?, \mathrm{~s}_{\mathrm{x}}, \mathrm{s}_{\mathrm{y}}$ ending at time $t$ is simply
$-\mathrm{P}\left(?, ?, ?, ?, \mathrm{~s}_{\mathrm{x}}, \mathrm{s}_{\mathrm{y}}\right)=\mathrm{P}\left(?, ?, ?, ?, \mathrm{~s}_{\mathrm{x}}\right) \mathrm{P}\left(o_{t} \mid s_{y}\right) \mathrm{P}\left(s_{y} \mid s_{x}\right)$
- The best state sequence that ends with $s_{x}, s_{y}$ at $t$ will have a probability equal to the probability of the best state sequence ending at $t-1$ at $s_{x}$ times $\mathrm{P}\left(o_{t} \mid s_{y}\right) \mathrm{P}\left(s_{y} \mid s_{x}\right)$
- Since the last term is independent of the state sequence leading to $s_{x}$ at $t-1$


## Trellis

- The graph below shows the set of all possible state sequences through this HMM in five time intants


The cost of extending a state sequence

- The cost of extending a state sequence ending at $s_{x}$ is only dependent on the transition from $s_{x}$ to $s_{y}$, and the observation probability at $s_{y}$


The cost of extending a state sequence

- The best path to $s_{y}$ through $s_{x}$ is simply an extension of the best path to $s_{x}$



## The Recursion

- The overall best path to $s_{x}$ is an extension of the best path to one of the states at the previous time



## The Recursion

- Bestpath prob $\left(\mathrm{s}_{\mathrm{y}}, \mathrm{t}\right)=$ Best $\left(\right.$ Bestpath $\left.\operatorname{prob}\left(\mathrm{s}_{?}, \mathrm{t}\right) * \mathrm{P}\left(\mathrm{s}_{\mathrm{y}} \mid \mathrm{s}_{?}\right) * \mathrm{P}\left(\mathrm{o}_{\mathrm{t}} \mid \mathrm{s}_{\mathrm{y}}\right)\right)$



## Finding the best state sequence

- This gives us a simple recursive formulation to find the overall best state sequence:

1. The best state sequence $X_{1, \mathrm{i}}$ of length 1 ending at state $s_{\mathrm{i}}$ is simply $s_{\mathrm{i}}$.

- The probability $\mathrm{C}\left(X_{1, \mathrm{i}}\right)$ of $X_{1, \mathrm{i}}$ is $P\left(o_{1} \mid s_{\mathrm{i}}\right) P\left(s_{\mathrm{i}}\right)$

2. The best state sequence of length $\mathrm{t}+1$ is simply given by
$-\quad\left(\operatorname{argmax}_{X_{\mathrm{t}, \mathrm{i}}} \mathrm{C}\left(X_{\mathrm{t}, \mathrm{i}}\right) P\left(o_{\mathrm{t}+1} \mid s_{\mathrm{j}}\right) P\left(s_{\mathrm{j}} \mid s_{\mathrm{i}}\right)\right) s_{\mathrm{i}}$
3. The best overall state sequence for an utterance of length $T$ is given by
$\operatorname{argmax}_{X_{\mathrm{t}, \mathrm{i}} s_{\mathrm{j}}} \mathrm{C}\left(X_{\mathrm{T}, \mathrm{i}}\right)$

- The state sequence of length T with the highest overall probability


## Finding the best state sequence

- The simple algorithm just presented is called the VITERBI algorithm in the literature
- After A.J.Viterbi, who invented this dynamic programming algorithm for a completely different purpose: decoding error correction codes!
- The Viterbi algorithm can also be viewed as a breadth-first graph search algorithm
- The HMM forms the Y axis of a 2-D plane
- Edge costs of this graph are transition probabilities $\mathrm{P}(\mathrm{s} \mid \mathrm{s})$. Node costs are $\mathrm{P}(\mathrm{o} \mid \mathrm{s})$
- A linear graph with every node at a time step forms the X axis
- A trellis is a graph formed as the crossproduct of these two graphs
- The Viterbi algorithm finds the best path through this graph


## Viterbi Search (contd.)




Initial state initialized with path-score $=P\left(s_{l}\right) b_{l}(l)$
time
${ }_{14 \text { Feb } 2011}$ All other states have score 0 since $P\left(s_{i}\right)=0$ for them

## Viterbi Search (contd.)



State with best path-score
State with path-score < best
State without a valid path-score

$$
P_{j}(t)=\max _{i}\left[P_{i}(t-1) a_{i j} b_{j}(t)\right]
$$

State transition probability, $i$ to $j$
Score for state $j$, given the input at time $t$
Total path-score ending up at state $j$ at time $t$

## Viterbi Search (contd.)



## Viterbi Search (contd.)


time

## Viterbi Search (contd.)



## Viterbi Search (contd.)


time

## Viterbi Search (contd.)



## Viterbi Search (contd.)



## Viterbi Search (contd.)

THE BEST STATE SEQUENCE IS THE ESTIMATE OF THE STATE SEQUENCE FOLLOWED IN GENERATING THE OBSERVATION


## Viterbi and DTW

- The Viterbi algorithm is identical to the stringmatching procedure used for DTW that we saw earlier
- It computes an estimate of the state sequence followed in producing the observation
- It also gives us the probability of the best state sequence


## Problem3: Training HMM parameters

- We can compute the probability of an observation, and the best state sequence given an observation, using the HMM's parameters
- But where do the HMM parameters come from?
- They must be learned from a collection of observation sequences
- We have already seen one technique for training HMMs: The segmental K-means procedure

Modified segmental K-means AKA
Viterbi training

- The entire segmental K-means algorithm:

1. Initialize all parameters

- State means and covariances
- Transition probabilities
- Initial state probabilities

2. Segment all training sequences
3. Reestimate parameters from segmented training sequences
4. If not converged, return to 2

## Segmental K-means



T1 T2 T3 T4
The procedure can be continued until convergence
Convergence is achieved when the total best-alignment error for all training sequences does not change significantly with further ${ }^{14}$ Feb zorefinement of the model

## A Better Technique

- The Segmental K-means technique uniquely assigns each observation to one state
- However, this is only an estimate and may be wrong
- A better approach is to take a "soft" decision
- Assign each observation to every state with a probability

Training by segmentation: Hard Assignment


## Training by segmentation: Soft Assignment


$i \in$ All vectors

$$
\sum_{\text {ell segments }} f_{i, j}=1
$$

Assignment is fractioned: Every segment gets a piece of every vector

Means and variances are computed from fractioned vectors

Where do the fractions come from?

## The "probability" of a state

- The probability assigned to any state $s$, for any observation $x_{t}$ is the probability that the process was at $s$ when it generated $x_{t}$
- We want to compute
$P\left(\right.$ state $\left.(t)=s \mid x_{1}, x_{2}, \ldots, x_{T}\right) \propto P\left(\right.$ state $\left.(t)=s, x_{1}, x_{2}, \ldots, x_{T}\right)$
- We will compute $P\left(\right.$ state $\left.(t)=s, x_{1}, x_{2}, \ldots, x_{T}\right)$ first
- This is the probability that the process visited $s$ at time $t$ while producing the entire observation

Probability of Assigning an Observation to a

## State

- The probability that the HMM was in a particular state $s$ when generating the observation sequence is the probability that it followed a state sequence that passed through $s$ at time $t$


Probability of Assigning an Observation to a State

- This can be decomposed into two multiplicative sections
- The section of the lattice leading into state $s$ at time $t$ and the section leading out of it


Probability of Assigning an Observation to a State

- The probability of the red section is the total probability of all state sequences ending at state $s$ at time $t$
- This is simply $\alpha(s, t)$
- Can be computed using the forward algorithm
time


## The forward algorithm


$\lambda$ represents the complete current set of HMM parameters

## The Future Paths

- The blue portion represents the probability of all state sequences that began at state $s$ at time $t$
- Like the red portion it can be computed using a backward recursion



## The Backward Recursion

$$
\begin{aligned}
& \beta_{u}(s, t)=P\left(x_{u, t+1}, x_{u, t+2}, \ldots, x_{u, T} \mid \text { state }(t)=s, \lambda\right) \\
& \bigcirc \bigcirc \\
& \begin{array}{ll} 
& \begin{array}{l}
\text { Can be recursively } \\
\text { estimated starting } \\
\text { from the final time }
\end{array} \\
\text { time instant }
\end{array} \\
& \beta_{u}(s, t)=\sum_{s^{\prime}} \beta_{u}\left(s^{\prime}, t+1\right) P\left(s^{\prime} \mid s\right) P\left(x_{u, t+1} \mid s^{\prime}\right)
\end{aligned}
$$

- $\beta_{u}(s, t)$ is the total probability of ALL state sequences that depart from $s$ at time $t$, and all observations after $x_{t}$
$-\beta(s, T)=1$ at the final time instant for all valid final states


## The complete probability

$$
\begin{aligned}
& \alpha_{u}(s, t) \beta_{u}(s, t)=P\left(x_{u, 1}, x_{u, 2}, \ldots, x_{u, T}, \operatorname{state}(t)=s \mid \lambda\right) \\
& =P\left(\mathbf{X}_{u}, \text { state }(t)=s \mid \lambda\right)
\end{aligned}
$$

## Posterior probability of a state

- The probability that the process was in state $s$ at time $t$, given that we have observed the data is obtained by simple normalization

$$
P\left(\text { state }(t)=s \mid \mathbf{X}_{u}, \lambda\right)=\frac{P\left(\mathbf{X}_{u}, \text { state }(t)=s \mid \lambda\right)}{\sum_{\gamma} P\left(\mathbf{X}_{u}, \text { state }(t)=s^{\prime} \mid \lambda\right)}=\frac{\alpha_{u}(s, t) \beta_{u}(s, t)}{\sum_{\gamma} \alpha_{u}\left(s^{\prime}, t\right) \beta_{u}\left(s^{\prime}, t\right)}
$$

- This term is often referred to as the gamma term and denoted by $\gamma_{\mathrm{s}, \mathrm{t}}$


## Update Rules

- Once we have the state probabilities (the gammas) the update rules are obtained through a simple modification of the formulae used for segmental K-means
- This new learning algorithm is known as the Baum-Welch learning procedure
- Case1: State output densities are Gaussians


## Update Rules

$$
\begin{gathered}
\mu_{s}=\frac{1}{N_{s}} \sum_{x \in s} x \\
C_{s}=\frac{1}{N_{s}} \sum_{x \in s}\left(x-\mu_{s}\right)^{T}\left(x-\mu_{s}\right)
\end{gathered} \begin{gathered}
\mu_{s}=\frac{\sum_{u} \sum_{t} \gamma_{u, s, t} x_{u, t}}{\sum_{u} \sum_{t} \gamma_{u, s, t}} \\
C_{s}=\frac{\sum_{u} \sum_{t} \gamma_{u, s, t}\left(x-\mu_{s}\right)^{T}\left(x-\mu_{s}\right)}{\sum_{u} \sum_{t} \gamma_{s, u, t}}
\end{gathered}
$$

Baum Welch

- A similar update formula reestimates transition probabilities
- The initial state probabilities $\mathrm{P}(s)$ also have a similar update rule

Case 2: State ouput densities are Gaussian

## Mixtures

- When state output densities are Gaussian mixtures, more parameters must be estimated

$$
P(x \mid s)=\sum_{i=0}^{K-1} w_{s, i} \operatorname{Gaussian}\left(x ; \mu_{s, i}, C_{s, i}\right)
$$

- The mixture weights $w_{s, i}$, mean $\mu_{s, i}$ and covariance $C_{s, i}$ of every Gaussian in the distribution of each state must be estimated


## Splitting the Gamma

We split the gamma for any state among all the Gaussians at that state


A posteriori probability that the $t^{\text {th }}$ vector was generated by the $\mathrm{k}^{\text {th }}$ Gaussian of state s

$$
\gamma_{k, s, u, t}=P\left(\operatorname{state}(t)=s \mid \mathbf{X}_{u}, \lambda\right) P\left(k^{\text {th }} . \operatorname{Gaussian} \mid \text { state }(t)=s, x_{u, t}, \lambda\right)
$$

## Splitting the Gamma among Gaussians

A posteriori probability that the $t^{\text {th }}$ vector was generated by the $\mathrm{k}^{\text {th }}$ Gaussian of state s

$$
\gamma_{k, s, u, t}=P\left(\operatorname{state}(t)=s \mid \mathbf{X}_{u}, \lambda\right) P\left(k^{t h} . \operatorname{Gaussian} \mid \operatorname{state}(t)=s, x_{u, t}, \lambda\right)
$$

$$
\gamma_{k, s, u, t}=P\left(\operatorname{state}(t)=s \mid \mathbf{X}_{u}, \lambda\right) \frac{w_{k, s} \frac{1}{\sqrt{(2 \pi)^{d}\left|\mathbf{C}_{k}\right|}} e^{-\frac{1}{2}\left(x_{u, t}-\mu_{k, s}\right)^{T} \mathbf{C}_{k}^{-1}\left(x_{u, t}-\mu_{k, s}\right)}}{\sum_{k^{\prime}} w_{k^{\prime}, s}} \frac{1}{\sqrt{(2 \pi)^{d}\left|\mathbf{C}_{k^{\prime}, s}\right|}} e^{-\frac{1}{2}\left(x_{u, t}-\mu_{k^{\prime}, s}\right)^{T} \mathbf{C}_{k}^{-1}\left(x_{u, t}-\mu_{k^{\prime}, s}\right)}
$$

## Updating HMM Parameters

$$
\begin{aligned}
& \tilde{\mu}_{k, s}=\frac{\sum_{u} \sum_{t} \gamma_{k, s, u, t} x_{u, t}}{\sum_{u} \sum_{t} \gamma_{k, s, u, t}} \widetilde{\mathbf{C}}_{k, s}=\frac{\sum_{u} \sum_{t} \gamma_{k, s, u, t}\left(x_{u, t}-\tilde{\mu}_{k, s}\right)\left(x_{u, t}-\tilde{\mu}_{k, s}\right)^{T}}{\sum_{u} \sum_{t} \gamma_{k, s, s, t}} \\
& \widetilde{w}_{k, s}=\frac{\sum_{u} \sum_{t} \gamma_{k, s, u, t}}{\sum_{u} \sum_{t} \sum_{j} \gamma_{j, s, u, t}}
\end{aligned}
$$

- Note: Every observation contributes to the update of parameter values of every Gaussian of every state

Overall Training Procedure: Single Gaussian

## PDF

- Determine a topology for the HMM
- Initialize all HMM parameters
- Initialize all allowed transitions to have the same probability
- Initialize all state output densities to be Gaussians
- We'll revisit initialization

1. Over all utterances, compute the "sufficient" statistics $\sum_{u} \sum_{t} \gamma_{u, s, t} \sum_{u} \sum_{t} \gamma_{u, s, t} x_{u, t} \sum_{u} \sum_{t} \gamma_{u, s, t}\left(x-\mu_{s}\right)^{T}\left(x-\mu_{s}\right)$
2. Use update formulae to compute new HMM parameters
3. If the overall probability of the training data has not converged, return to step 1

## An Implementational Detail

- Step1 computes "buffers" over all utterance

$$
\sum_{u} \sum_{t} \gamma_{u, s, t} x_{u, t}=\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t} x_{u, t}+\sum_{u \in U_{2}} \sum_{t} \gamma_{u, s, t} x_{u, t}+\ldots
$$



- This can be split and parallelized
$-\mathrm{U}_{1}, \mathrm{U}_{2}$ etc. can be processed on separate machines

$$
\begin{gathered}
\text { Machine 1 } \\
\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t} \sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t} x_{u, t} \\
\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t}\left(x-\mu_{s}\right)^{T}\left(x-\mu_{s}\right)
\end{gathered}
$$

## An Implementational Detail

- Step2 aggregates and adds buffers before updating the models

$$
\sum_{u} \sum_{t} \gamma_{u, s, t}=\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t}+\sum_{u \in U_{2}} \sum_{t} \gamma_{u, s, t}+.
$$

$$
\sum_{u} \sum_{t} \gamma_{u, s, t} x_{u, t}=\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t} x_{u, t}+\sum_{u \in U_{2}} \sum_{t} \gamma_{u, s, t} x_{u, t}+\ldots
$$



$$
\tilde{\mu}_{k, s}=\frac{\sum_{u} \sum_{t} \gamma_{k, s, u, t} x_{u, t}}{\sum_{u} \sum_{t} \gamma_{k, s, u, t}} \widetilde{\mathbf{C}}_{k, s}=\frac{\sum_{u} \sum_{t} \gamma_{k, s, u, t}\left(x_{u, t}-\tilde{\mu}_{k, s}\right)\left(x_{u, t}-\tilde{\mu}_{k, s}\right)^{T}}{\sum_{u} \sum_{t} \gamma_{k, s, u, t}}
$$

$$
\widetilde{w}_{k, s}=\frac{\sum_{u} \sum_{t} \gamma_{k, s, u, t}}{\sum_{u} \sum_{t} \sum_{j} \gamma_{j, s, u, t}}
$$

## An Implementational Detail

- Step2 aggregates and adds buffers before updating the models

$$
\begin{aligned}
& \sum_{u} \sum_{t} \gamma_{u, s, t}=\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t}+\sum_{u \in U_{2}} \sum_{t} \gamma_{u, s, t}+\ldots \\
& \sum_{u} \sum_{t} \gamma_{u, s, t} x_{u, t}=\sum_{u \in U_{1}} \sum_{t} \gamma_{u, s, t} x_{u, t}+\sum_{u \in U_{2}} \sum_{t} \gamma_{u, s, t} x_{u, t}+\ldots
\end{aligned}
$$

$$
\begin{aligned}
& \widetilde{w}_{k, s}=\frac{\sum_{u} \sum_{1} \gamma_{k, \ldots, t}}{\sum_{u} \sum_{1} \sum_{1} \gamma_{j, s, u t}} \\
& \text { Computed by } \\
& \text { machine } 1 \\
& \text { Computed by } \\
& \text { machine } 2
\end{aligned}
$$

Training for HMMs with Gaussian Mixture

## State Output Distributions

- Gaussian Mixtures are obtained by splitting

1. Train an HMM with (single) Gaussian state output distributions
2. Split the Gaussian with the largest variance

- Perturb the mean by adding and subtracting a small number
- This gives us 2 Gaussians. Partition the mixture weight of the Gaussian into two halves, one for each Gaussian
- A mixture with N Gaussians now becomes a mixture of $\mathrm{N}+1$ Gaussians

3. Iterate BW to convergence
4. If the desired number of Gaussians not obtained, return to 2

## Splitting a Gaussian



- The mixture weight $w$ for the Gaussian gets shared as $0.5 w$ by each of the two split Gaussians


## Transition Probabilities



- We have seen how to compute transition penalties for templates

- How about transition probabilities in an HMM?
- "Hard" estimation - by counting, as for templates
- "Soft" estimation - need soft counts


## Transition penalties by counting



- 20 vectors in state 1
- 16 are followed by vectors in state 1
- 4 are followed by vectors in state 2
- $\mathrm{P}_{11}=16 / 20=0.8 \rightarrow \mathrm{~T}_{11}=-\log \left(\mathrm{P}_{11}\right)=-\log (0.8)$
- $\mathrm{P}_{12}=4 / 20=0.2 \rightarrow \mathrm{~T}_{12}=-\log \left(\mathrm{P}_{12}\right)=-\log (0.2)$


## Transitions by counting



- We found the best state sequence for each input
- And counted transitions


## Transitions by counting

Observation no. 6 contributed one to the count of occurrences of state 1 Observations 6 and 7 contributed one to the count of transitions from state 1 to state 2


- We found the best state sequence for each input
- And counted transitions


## Probability of transitions



- $\mathrm{P}($ transition state $\mathrm{I} \rightarrow$ state $\mathrm{J}=$
- Count transitions(I,J) / count instances(I)


## Probability of transitions



- $\mathrm{P}($ transition state $\mathrm{I} \rightarrow$ state $\mathrm{J}=$
- Count transitions(I,J) / count instances(I)
- Count instances $(1)=20$


## Probability of transitions



- $\mathrm{P}($ transition state $\mathrm{I} \rightarrow$ state $\mathrm{J}=$
- Count transitions(I,J) / count instances(I)
- Count instances $(1)=20$
- Count transitions $(1,1)=16$
- $\mathrm{P}($ transition state $1 \rightarrow$ state 1$)=0.8$


## Probability of transitions



- $\mathrm{P}($ transition state $\mathrm{I} \rightarrow$ state $\mathrm{J}=$
- Count transitions(I,J) / count instances(I)
- Count instances(1) $=20$
- Count transitions $(1,2)=4$
- $\mathrm{P}($ transition state $1 \rightarrow$ state 2$)=0.2$


## Transitions by soft counting



- Each observation pair contributes to every transition
- E.g. observations 6,7 contribute counts to all of the following:
- Transition $(1 \rightarrow 1)$, Transition $(1 \rightarrow 2)$, Transition $(2 \rightarrow 2)$, Transition $(2 \rightarrow 3)$, Transition $(3 \rightarrow 3)$


## Transitions by soft counting



- Contribution of any transition to the count is the a posteriori probability of the count
- This is a fraction
- The fractions for all possible transitions at any time sum to 1


## Transitions by soft counting



- Probability of a transition is the total probability of all paths that include the transition


## Transitions by soft counting



- The forward probability of the source state at $t$ accounts for all incoming paths at time $t$
- including the $t$-th observation $x_{t}$


## Transitions by soft counting



- The backward probability of the destination state at $t+1$ accounts for all outgoing paths from the state at time $t+1$
- NOT including the $t+1$-th observation $x_{t+1}$


## Transitions by soft counting



- The product of the forward probability of $s$ at $t$ and $s^{\prime}$ at $t+1$ accounts for all paths TO state $s$ at $t$, and all paths FROM $s^{\prime}$ at $t+1$
- But not the transition from $s$ to $s$ ' or the observation at $t+1$


## Transitions by soft counting



- By factoring in the transition probability and observation probabilities, the total probability is obtained


## From probability to a posteriori probability



- The a posteriori probability of a transition is the ratio of its probability to the sum of all transitions at the same time


## A posteriori probability of a transition

- Probability of a transition
$P\left(\operatorname{state}(t)=s, \operatorname{state}(t+1)=s^{\prime}, x_{1}, x_{2}, \ldots \ldots x_{N}\right)=\alpha_{u}(s, t) P\left(s^{\prime} \mid s\right) P\left(x_{t}+1 \mid s^{\prime}\right) \beta_{u}\left(s^{\prime}, t+1\right)$
- A posteriori probability of a transition

$$
\gamma_{u, s, t, s^{\prime}, t+1}=\frac{P\left(\operatorname{state}(t)=s, \text { state }(t+1)=s^{\prime}, x_{1}, x_{2}, \ldots x_{N}\right)}{\sum_{S, S^{\prime}} P\left(\operatorname{state}(t)=S, \operatorname{state}(t+1)=S^{\prime}, x_{1}, x_{2}, \ldots x_{N}\right)}
$$

## Estimate of Transition Probabilities

$$
P\left(s^{\prime} \mid s\right)=\frac{\sum_{u} \sum_{t} \gamma_{u, s, t, s^{\prime}, t+1}}{\sum_{u} \sum_{t} \gamma_{u, s, t}}
$$

- Numberator is total "soft" count of transitions from state $s$ to $s$,
- Denumberator is total "soft" count of instances of state $s$


## Implementation of BW: underflow

- Arithmetic underflow is a problem

- The alpha terms are a recursive product of probability terms
- As tincreases, an increasingly greater number probability terms are factored into the alpha
- All probability terms are less than 1
- State output probabilities are actually probability densities
- Probability density values can be greater than 1
- On the other hand, for large dimensional data, probability density values are usually much less than 1
- With increasing time, alpha values decrease
- Within a few time instants, they underflow to 0
- Every alpha goes to 0 at some time $t$. All future alphas remain 0
- As the dimensionality of the data increases, alphas goes to 0 faster


## Underflow: Solution

- One method of avoiding underflow is to scale all alphas at each time instant
- Scale with respect to the largest alpha to make sure the largest scaled alpha is 1.0
- Scale with respect to the sum of the alphas to ensure that all alphas sum to 1.0
- Scaling constants must be appropriately considered when computing the final probabilities of an observation sequence
- An alternate method: Compute alphas and betas in log domain
- How? (Not obvious)


## Implementation of BW: underflow

- Similarly, arithmetic underflow can occur during beta computation

$$
\beta_{u}(s, t)=\sum_{s^{\prime}} \beta_{u}\left(s^{\prime}, t+1\right) \log P\left(s^{\prime} \mid s\right) P\left(x_{u, t+1} \mid s^{\prime}\right)
$$

- The beta terms are also a recursive product of probability terms and can underflow
- Underflow can be prevented by
- Scaling: Divide all beta terms by a constant that prevents underflow
- By performing beta computation in the log domain (now? Not obvious..)
- QUESTION: HOW DOES SCALING AFFECT THE ESTIMATION OF GAMMA TERMS
- For Gaussian parameter updates?
- For transition probability updates?


## Implementation of BW: pruning


= pruned out

- The forward backward computation can get very expensive
- Solution: Prune
- Pruning in the forward backward algorithm raises some additional issues
- Pruning from forward pass can be employed by backward pass
- Convergence criteria and tests may be affected
- More later


## Building a recognizer for isolated words

- Now have all necessary components to build an HMM-based recognizer for isolated words
- Where each word is spoken by itself in isolation
- E.g. a simple application, where one may either say "Yes" or "No" to a recognizer and it must recognize what was said


## Isolated Word Recognition with HMMs

- Assuming all words are equally likely
- Training
- Collect a set of "training" recordings for each word
- Compute feature vector sequences for the words
- Train HMMs for each word
- Recognition:
- Compute feature vector sequence for test utterance
- Compute the forward probability of the feature vector sequence from the HMM for each word
- Alternately compute the best state sequence probability using Viterbi
- Select the word for which this value is highest


## Issues

- What is the topology to use for the HMMs
- How many states
- What kind of transition structure
- If state output densities have Gaussian Mixtures: how many Gaussians?


## HMM Topology

- For speech a left-to-right topology works best
- The "Bakis" topology
- Note that the initial state probability $\mathrm{P}(\mathrm{s})$ is 1 for the $1^{\text {st }}$ state and 0 for others. This need not be learned

- States may be skipped



## Determining the Number of States

- How do we know the number of states to use for any word?
- We do not, really
- Ideally there should be at least one state for each "basic sound" within the word
- Otherwise widely differing sounds may be collapsed into one state
- The average feature vector for that state would be a poor representation
- For computational efficiency, the number of states should be small
- These two are conflicting requirements, usually solved by making some educated guesses


## Determining the Number of States

- For small vocabularies, it is possible to examine each word in detail and arrive at reasonable numbers:

$\begin{array}{llllll}S & O & M E & T H & I & N G\end{array}$
- For larger vocabularies, we may be forced to rely on some $a d$ hoc principles
- E.g. proportional to the number of letters in the word
- Works better for some languages than others
- Spanish and Indian languages are good examples where this works as almost every letter in a word produces a sound


## How many Gaussians

- No clear answer for this either
- The number of Gaussians is usually a function of the amount of training data available
- Often set by trial and error
- A minimum of 4 Gaussians is usually required for reasonable recognition

Implementation of BW: initialization of alphas and betas

- Initialization for alpha: $\alpha_{u}(s, 1)$ set to 0 for all states except the first state of the model. $\alpha_{u}(s, 1)$ set to 1 for the first state
- All observations must begin at the first state
- Initialization for beta: $\beta_{u}(s, T)$ set to 0 for all states except the terminating state. $\beta_{u}(s, t)$ set to 1 for this state
- All observations must terminate at the final state


## Initializing State Output Density Parameters

1. Initially only a single Gaussian per state assumed

- Mixtures obtained by splitting Gaussians

2. For Bakis-topology HMMs, a good initialization is the "flat" initialization

- Compute the global mean and variance of all feature vectors in all training instances of the word
- Initialize all Gaussians (i.e all state output distributions) with this mean and variance
- Their means and variances will converge to appropriate values automatically with iteration
- Gaussian splitting to compute Gaussian mixtures takes care of the rest


## Isolated word recognition: Final thoughts

- All relevant topics covered
- How to compute features from recordings of the words
- We will not explicitly refer to feature computation in future lectures
- How to set HMM topologies for the words
- How to train HMMs for the words
- Baum-Welch algorithm
- How to select the most probable HMM for a test instance
- Computing probabilities using the forward algorithm
- Computing probabilities using the Viterbi algorithm
- Which also gives the state segmentation


## Questions

- ?

